Model Selection and Variable Selection

Bayesian Methodology in Biostatistics (BST 249)

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Outline

Introduction

Variable selection

Oxygen update example Diabetes example

Bayesian variable selection

Oxygen uptake example, revisited

Gibbs sampling for Bayesian variable selection Diabetes example, revisited

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Introduction

- When using regression in practice, many possible variables could be included as covariates.
- Including too many variables can lead to poor prediction performance.
- Ideally, one would include only the variables that are informative about the outcome.
- But how can one determine which variables are informative?
- This is the variable selection problem.
- Variable selection is a particular type of model selection.

Bayesian model selection

- The standard Bayesian approach to model selection is to put a prior on models, and simply consider the posterior on models.
- This is sometimes referred to as "Bayesian model averaging", rather than "model selection".
- The "model averaging" terminology emphasizes the fact that we use the posterior over all models rather than selecting a single model.
- Thus, posterior expectations involve averaging over models.

Bayesian model selection

- Suppose we are considering models $\mathcal{M}_1, \mathcal{M}_2, \ldots$
- Suppose model \mathcal{M}_k has parameter θ_k .
- Suppose we have a prior on the model index p(k), and a prior on parameters $p(\theta_k \mid k)$ for each model.
- Given data x, the posterior on models is then

$$p(k|x) \propto p(x|k) p(k)$$

= $p(k) \int p(x \mid \theta_k, k) p(\theta_k \mid k) d\theta_k.$

• p(x|k) is the marginal likelihood of model k. Incidentally, this justifies the term "marginal likelihood", since it is the likelihood function for k.

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Variable selection

- Variable selection is a special case of model selection.
- Consider each subset of possible variables to be a different model.
- Represent this as follows: $z_j = 1$ if variable j is included, and $z_j = 0$ otherwise.
- Consider the linear regression model:

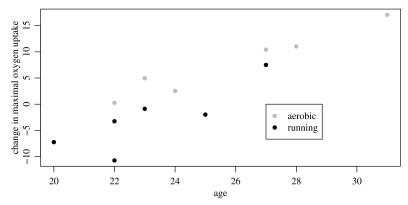
$$Y_i = z_1 \beta_1 x_{i1} + \dots + z_p \beta_p x_{ip} + \varepsilon_i.$$

- Each possible vector $z = (z_1, \ldots, z_p) \in \{0, 1\}^p$ represents a different model.
- So, for model selection, we are interested in $p(z \mid x, y)$.

Oxygen update example: Data

- Twelve healthy men were recruited to take part in a study on the effects of exercise. The men did not exercise regularly.
- Subjects were randomly divided into two groups of six.
 - Group 1 followed a 12-week running program.
 - Group 2 followed a 12-week step aerobics program.
- Maximum oxygen uptake (liters/minute) was measured while running on a treadmill, both before and after the program.
- Goal: Assess the effect of the running program on oxygen uptake.

Oxygen update example: Data



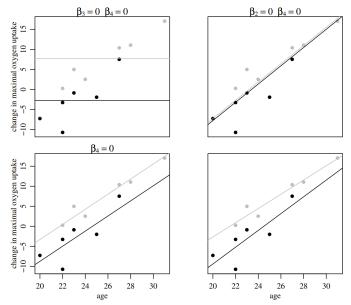
Change in maximal oxygen uptake as a function of age and exercise program.

Oxygen update example: Models considered

• Five models under consideration:

 $E(Y \mid X, \beta, z = (1, 0, 0, 0)) = \beta_1$ $E(Y \mid X, \beta, z = (1, 1, 0, 0)) = \beta_1 + \beta_2$ group $E(Y \mid X, \beta, z = (1, 0, 1, 0)) = \beta_1 + \beta_3$ age $E(Y \mid X, \beta, z = (1, 1, 1, 0)) = \beta_1 + \beta_2 \operatorname{group} + \beta_3 \operatorname{age}$ $E(Y \mid X, \beta, z = (1, 1, 1, 1))$ $=\beta_1+\beta_2$ group $+\beta_3$ age $+\beta_4$ group \times age.

Oxygen update example: OLS for four models



Least squares regression lines for the oxygen uptake data, under four different models.

Oxygen update example: Motivation

- Visually, some of these models clearly seem better than others.
- Statistically, how much evidence is there for each model?
- A frequentist approach would employ hypothesis testing, but this gets complicated when there are many possible variables to include.
- It is common to use lasso, elastic net, or stepwise selection, but significance testing is not simple with these methods. Recent methods for "post-selection inference" have been developed to address this.
- The Bayesian approach is simply to consider the posterior on which variables to include.

Diabetes example: Data

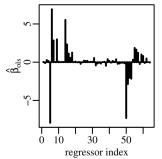
- n = 442 patients, 10 binary measurements for each patient.
- Outcome: y_i = quantitative measure of disease progression of patient *i*, one year after measurements.
- Goal: Predict y_i from the measurements.
- Baseline model: Linear regression with main effects as well as interactions between each pair of measurements.
- The baseline model has p = 64 covariates, which are centered and scaled for interpretability.
- 342 patients used for training, 100 patients used for testing.

Diabetes example: Performance comparison

- Approach #1: Constant prediction $\hat{y}_i = c$. Test MSE = 0.97.
- Approach #2: OLS including all covariates. Test MSE = 0.67.
- Approach #3: OLS + backward selection. Test MSE = 0.53.
- Approach #4: Bayesian variable selection. Test MSE = 0.45.

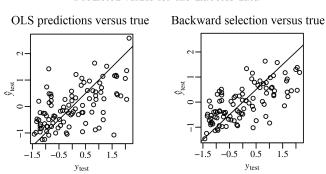
Diabetes example: OLS coefficients

OLS coefficients for diabetes data



(figure from Hoff, 2009)

Diabetes example: OLS versus backward selection

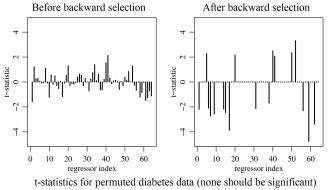


Predicted values for the diabetes data

(figure from Hoff, 2009)

- Backward selection starts with all variables included, and iteratively removes the least significant variable and refits, until all variables are significant. This leaves 20 variables.
- Backward selection helps improve prediction performance.

Diabetes example: Issue with backward selection



(figure from Hoff, 2009)

- Unfortunately, backward selection often indicates relationships even when there are none.
- Above, the *y* values were randomly permuted, breaking any dependencies with the *x*'s.
- Backward selection incorrectly finds many "significant" coefficients.

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Bayesian variable selection

• Consider the linear regression model:

$$Y_i = z_1\beta_1 x_{i1} + \dots + z_p\beta_p x_{ip} + \varepsilon_i$$
 where $\varepsilon_i \sim \mathcal{N}(0,\gamma^{-1}).$

- Prior on models: $z \sim p(z)$, e.g., $Z_1, \ldots, Z_p \stackrel{\text{iid}}{\sim} \text{Bernoulli}(\alpha)$.
- Prior on parameters for each model:

$$\gamma \sim \text{Gamma}(\frac{1}{2}\nu_0, \frac{1}{2}\nu_0\sigma_0^2)$$

$$\beta_z | X, z, \gamma \sim \mathcal{N}(0, g(\gamma X_z^{\mathsf{T}} X_z)^{-1})$$

where $\beta_z = (\beta_j : z_j = 1)$ contains the entries of β where $z_j = 1$, and likewise, X_z is the design matrix including only the columns j for which $z_j = 1$. If $z_j = 0$, then set β_j to any arbitrary real value, e.g., $\beta_j = 0$.

• Thus, given z, this is a g-prior on β_z .

Bayesian variable selection: Posterior computation

• The posterior on models z, integrating out β and $\gamma,$ is

 $p(z|X,y) \propto p(y|X,z)p(z).$

• It turns out that p(y|X,z) can be computed analytically (see Hoff 9.3.1):

$$p(y|X,z) = \int p(y \mid \beta, \gamma, X, z) p(\beta \mid \gamma, X, z) p(\gamma) \, d\beta \, d\gamma \quad (1)$$
$$= \frac{1}{\pi^{n/2} (1+g)^{p_z/2}} \frac{\Gamma(\frac{1}{2}(\nu_0 + n))}{\Gamma(\frac{1}{2}\nu_0)} \frac{(\nu_0 \sigma_0^2)^{\nu_0/2}}{(\nu_0 \sigma_0^2 + \mathrm{SSR}_g^z)^{(\nu_0 + n)/2}}$$

where $p_z = \sum_{j=1}^p z_j$ and

$$\mathrm{SSR}_g^z = y^{\mathrm{T}} \Big(I - \frac{g}{g+1} X_z (X_z^{\mathrm{T}} X_z)^{-1} X_z^{\mathrm{T}} \Big) y.$$

Bayesian variable selection: Prior settings

- A natural default is to use a unit information prior.
- Unit information prior: g = n, $\nu_0 = 1$, and $\sigma_0^2 = \hat{\sigma}_{mle}^2$.
- For the prior on models, in these examples, we will consider a uniform prior $p(z) \propto 1$. Equivalently, $Z_j \stackrel{\text{iid}}{\sim} \text{Bernoulli}(1/2)$.
- However, to favor sparsity, it is common to use $Z_j \stackrel{\text{iid}}{\sim} \text{Bernoulli}(\alpha)$ where α is of order 1/p.
- For instance, if $\alpha = c/p$ then (under the prior) the proportion of included coefficients is c.
- It is also common (and recommended) to integrate out a Beta prior on α, which can easily be done analytically since it is conjugate.

Oxygen uptake example: Posterior on models

Marginal probabilities of the data under five different models.

\boldsymbol{z}	model	$\log p(\boldsymbol{y} \mathbf{X}, \boldsymbol{z})$	$p(\boldsymbol{z} \boldsymbol{y}, \mathbf{X})$
(1,0,0,0)	β_1	-44.33	0.00
(1,1,0,0)	$\beta_1 + \beta_2 \times \operatorname{group}_i$	-42.35	0.00
(1,0,1,0)	$\beta_1 + \beta_3 \times \text{age}_i$	-37.66	0.18
(1,1,1,0)	$\beta_1 + \beta_2 \times \operatorname{group}_i + \beta_3 \times \operatorname{age}_i$	-36.42	0.63
	$\beta_1 + \beta_2 \times \operatorname{group}_i + \beta_3 \times \operatorname{age}_i + \beta_4 \times \operatorname{group}_i \times \operatorname{age}_i$	-37.60	0.19

(figure from Hoff, 2009)

- The posterior favors the model including age and the exercise group, but not the interaction.
- However, a sizable amount of posterior mass is also given to the other two models that include age.
- Thus, according to this analysis, the data provides some evidence that the type of exercise program has an effect, but it is not definitive.

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- In the oxygen uptake example, there are only 5 models under consideration, so we can easily analytically compute the posterior over all 5 models.
- However, when there are more variables, the number of models will be far too large to consider them all.
- If we consider all subsets of p variables, there are 2^p possible models.
- For instance, for the diabetes data, p=64, so there are around 1.8×10^{19} models!
- Gibbs sampling is a common approach to doing approximate posterior inference for variable selection.

- As mentioned earlier, one can integrate out β and γ to obtain an analytic expression for p(y|X, z).
- We can use this expression to do Gibbs sampling directly on p(z|X,y), since $p(z|X,y) \propto p(y|X,z)p(z)$.
- Gibbs sampler algorithm for variable selection
 - lnitialize $z_1 = \cdots = z_p = 0.$
 - At each iteration, for each j = 1, ..., JUpdate z_j by sampling from $p(z_j | X, y, z_{-j})$.
- Here, z_{-j} denotes all the entries of z except z_j .
- Note: Initializing $z_j = 0$ speeds up burn-in when p is very large and the posterior is concentrated on sparse z vectors.

• The full conditional for z_j is

$$p(z_j \mid X, y, z_{-j}) \underset{z_j}{\propto} p(y \mid X, z) p(z)$$

and p(y|X, z) is given by Equation 1 (earlier in these slides).

• This can be written as

$$p(z_j = 1 \mid X, y, z_{-j}) = r_j / (1 + r_j)$$

$$p(z_j = 0 \mid X, y, z_{-j}) = 1 / (1 + r_j)$$

where

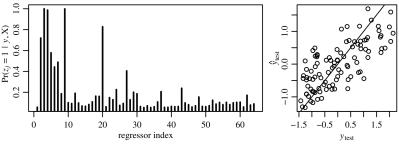
$$r_j = r_j(z_{-j}) = \frac{p(y \mid X, z_{-j}, z_j = 1)}{p(y \mid X, z_{-j}, z_j = 0)} \frac{p(z_{-j}, z_j = 1)}{p(z_{-j}, z_j = 0)}.$$

- A Gibbs sampler in which some of the variables have been integrated out is sometimes called a "collapsed Gibbs sampler".
- What if we want to get posterior samples of β and γ as well? It turns out that this is easy to do.
- First run the Gibbs sampler for T iterations to get samples $z^{(1)},\ldots,z^{(T)}$ from p(z|X,y).
- Then, for each $t = 1, \dots, T$, 1. Sample $\gamma^{(t)} \sim p(\gamma \mid X, y, z^{(t)})$. 2. Sample $\beta^{(t)} \sim p(\beta \mid X, y, z^{(t)}, \gamma^{(t)})$.
- Sampling $\gamma^{(t)}$ and $\beta^{(t)}$ can be done exactly, using the formulas from the slides on g-priors for Bayesian linear regression.

Diabetes example: Posterior

- Run sampler for T=10000 iterations, to get $z^{(t)},\gamma^{(t)},\beta^{(t)}$ for $t=1,\ldots,T.$
- Only a small fraction of the total number of models are ever visited by the sampler.
- However, if the sampler is performing reasonably well, then the set of models that are not visited should have small posterior probability.
- Further, the samples often provide a reasonable approximation to the marginal distribution of each z_j and β_j, even if the joint posterior on z, β is not very well approximated.

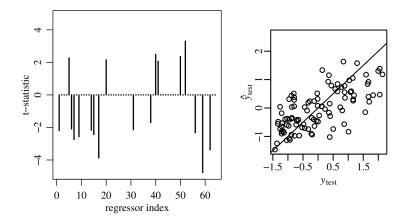
Diabetes example: Posterior



The first panel shows posterior probabilities that each coefficient is non-zero. The second panel shows y_{test} versus predictions based on the model averaged estimate of β .

- Test MSE = 0.45 for Bayesian variable selection.
- Bayes selects quite different variables than backward selection.

Diabetes example: Compare with backward selection



• Test MSE = 0.53 for backward selection.

References and supplements

- Hoff, P. D. (2009). A first course in Bayesian statistical methods (Vol. 580). New York: Springer.
- George, E. I., & McCulloch, R. E. (1993). Variable selection via Gibbs sampling. Journal of the American Statistical Association, 88(423), 881-889.