## Posterior consistency for

 the number of components in a finite mixtureJeffrey W. Miller<br>and<br>Matthew T. Harrison

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## Summary

1 Dirichlet process mixtures (DPMs) are not consistent for the number of components in a finite mixture.

2 However, there is a natural alternative that is consistent and exhibits many of the attractive properties of DPMs.

## Dirichlet process mixture (DPM) model

## Generic DPM model

$Q \sim \mathrm{DP}(\alpha, H)$
$\beta_{1}, \beta_{2}, \ldots \stackrel{\text { iid }}{\sim} Q$ (given $Q$ )
$X_{i} \sim p_{\beta_{i}}$ independent for $i=1,2, \ldots$ (given $Q, \beta_{1}, \beta_{2}, \ldots$ ) for some parametric family $\left\{p_{\theta}: \theta \in \Theta\right\}$.


Let $T_{n}=\#\left\{\beta_{1}, \ldots, \beta_{n}\right\}$.
That is, $T_{n}$ is the number of distinct components so far (i.e. the number of clusters).

## Mixture of finite mixtures (MFM)

Many authors have considered the following natural alternative to DPMs.
e.g. Nobile (1994, 2000, 2004, 2005, 2007), Richardson \& Green (1997, 2001), Stephens (2000),

Zhang et al. (2004), Kruijer (2008), Rousseau (2010), Kruijer, Rousseau, \& van der Vaart (2010).
Instead of $Q \sim \mathrm{DP}(\alpha, H)$, choose $Q$ as follows:
A mixture over finite mixtures
$S \sim p(s)$, a p.m.f. on $\{1,2, \ldots\}$
$\pi \sim \operatorname{Dirichlet}\left(\gamma_{s 1}, \ldots, \gamma_{s s}\right)$ (given $S=s$ )
$\theta_{1}, \ldots, \theta_{s} \stackrel{\text { iid }}{\sim} H$ (given $S=s$ )
$Q=\sum_{i=1}^{S} \pi_{i} \delta_{\theta_{i}}$


For mathematical convenience, we suggest:

- $H$ as a conjugate prior for $\left\{p_{\theta}\right\}$
- $p(s)=\operatorname{Poisson}(s-1 \mid \lambda)$
- $\gamma_{i j}=\gamma>0$ for all $i, j$


## Questions of convergence

## For data from a finite mixture, is the posterior consistent ...

(and at what rate of convergence) ...

|  | DPMs | MFMs |
| :--- | :--- | :--- |
| $\ldots$ for the density? | Yes (optimal rate) | Yes (optimal rate) |

DPMs: Ghosal \& van der Vaart $(2001,2007)$, and others.
MFMs: Doob's theorem gives a.e. consistency. Kruijer et al. $(2008,2010)$ prove rates.
...for the mixing distribution? Yes (optimal rate) Yes
DPMs: Nguyen (2012)
MFMs: Doob's theorem gives a.e. consistency. Optimal rate?
...for the number of components? Not consistent Yes
DPMs: This is our contribution.
MFMs: Doob's theorem gives a.e. consistency (see e.g. Nobile (1994)).
(Note: Ignoring tiny clusters might fix this issue.)

## Toy example \#1: One normal component

Prior ( x ) and estimated posterior ( o ) of the number of clusters









Data: $\mathcal{N}(0,1)$. Each plot is the average over 5 datasets. Burn-in: 10,000 sweeps, Sample: 100,000 sweeps.
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## Toy example \#2: Two normal components



Data: $\frac{1}{2} \mathcal{N}(0,1)+\frac{1}{2} \mathcal{N}(6,1)$. Each plot is the average over 5 datasets. Burn-in: 10,000 sweeps, Sample: 100,000 sweeps.

## Toy example \#3: Five normal components

Prior ( x ) and estimated posterior ( o ) of the number of clusters









Data: $\sum_{k=-2}^{2} \frac{1}{5} \mathcal{N}\left(4 k, \frac{1}{2}\right)$. Each plot is the average over 5 datasets. Burn-in: 10,000 sweeps, Sample: 100,000 sweeps.

## The wrong intuition

It is tempting to think that the prior on the number of clusters $T_{n}$ is the culprit. After all, when e.g. $\alpha=1$,

$$
P_{\mathrm{DPM}}\left(T_{n}=t\right)=\frac{1}{n!}\left[\begin{array}{l}
n \\
t
\end{array}\right] \sim \frac{1}{n} \frac{(\log n)^{t-1}}{(t-1)!}=\operatorname{Poisson}(t-1 \mid \log n)
$$

where $\left[\begin{array}{l}n \\ t\end{array}\right]$ is an (unsigned) Stirling number of the first kind, and $a_{n} \sim b_{n}$ means that $a_{n} / b_{n} \rightarrow 1$ as $n \rightarrow \infty$. Hence, $P_{\mathrm{DPM}}\left(T_{n}=t\right) \rightarrow 0$ for any $t$.


However, this is not the fundamental reason why inconsistency occurs. Even if we replace the prior on $T_{n}$ by something that is not diverging, inconsistency remains!

## Comparing DPMs to MFMs

## Similarities between DPMs to MFMs:

- Efficient approximate inference (via Gibbs sampling)
- Appealing equivalent formulations:
- exchangeable distribution on partitions. . e.g. when $\alpha=1$ and $\gamma=1$ :

$$
P_{\mathrm{DPM}}(\mathcal{C})=\frac{1}{n!} \prod_{c \in \mathcal{C}}(|c|-1)!\quad \text { and } \quad P_{\mathrm{MFM}}(\mathcal{C})=\kappa(n, t) \prod_{c \in \mathcal{C}}|c|!
$$

- restaurant process
- stick-breaking
- random discrete measures
- Consistent at any sufficiently smooth density (at optimal rate, in a certain sense)

Advantages of MFMs (for data from a finite mixture):

- MFMs are a natural Bayesian extension of finite mixtures
- Consistency (a.e.) for $S, \pi, \theta$, and the density is automatically guaranteed


## Disadvantages of MFMs:

- More parameters (. . you have to choose $p(s)$ )
- (Slightly) more complicated sampling formulas


# Thank you! 

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